Mean-Field fluctuations in Quadratic interaction models in low SNR regimes

Nabarun Deb Statistics and Probability Seminar Department of Mathematics and Statistics Boston University

Joint work with Seunghyun (Sky) Li and Sumit Mukherjee

Outline

- Introduction
- 2 The Naive Mean-Field method
- 3 Two-spin Ising model: LLN + CLT
- 4 Bayesian Linear regression : LLN + CLT
- 6 Conclusion

Ising Model

Consider

$$\frac{d\mathbb{P}}{d\prod_{i=1}^{n}\mu_{i}}(\boldsymbol{\beta}) := \frac{1}{Z_{n}} \exp\left(\frac{1}{2}\boldsymbol{\beta}^{\top}\mathbf{A}_{n}\boldsymbol{\beta} + \mathbf{c}^{\top}\boldsymbol{\beta}\right),$$

$$\mu_i$$
 supported on $[-1,1]$, $\boldsymbol{\beta} = (\beta_1, \dots, \beta_n)$, and field vector $\mathbf{c} = (c_1, \dots, c_n)$.

Ising Model

Consider

$$\frac{d\mathbb{P}}{d\prod_{i=1}^{n}\mu_{i}}(\boldsymbol{\beta}) := \frac{1}{Z_{n}} \exp\left(\frac{1}{2}\boldsymbol{\beta}^{\top} \mathbf{A}_{n} \boldsymbol{\beta} + \mathbf{c}^{\top} \boldsymbol{\beta}\right),$$

 μ_i supported on [-1,1], $\boldsymbol{\beta} = (\beta_1, \dots, \beta_n)$, and field vector $\mathbf{c} = (c_1, \dots, c_n)$.

- **1** $\mathbf{A}_n = 0$ implies β_i s are independent
- **2** $\mathbf{A}_n(i,j) > 0$ implies that sites i and j are inclined to align in the same direction.
- lacktriangle Large c_i implies site i is more likely to take larger values
- **1** Interaction matrix \mathbf{A}_n , Partition function Z_n

Our goal

To study the asymptotic distribution of

$$T_n = \mathbf{q}^{\top} (\boldsymbol{\beta} - ??), \quad \|\mathbf{q}\| = 1$$

for certain linear combinations.

$$\mathbb{P}(oldsymbol{eta}) := rac{1}{Z_n} \exp\left(rac{ heta}{2} oldsymbol{eta}^ op \mathbf{A}_n oldsymbol{eta} + B \sum_{i=1}^n eta_i
ight),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is constant at B. θ — temperature parameter.

$$\mathbb{P}(\boldsymbol{\beta}) := \frac{1}{Z_n} \exp\left(\frac{\theta}{2} \boldsymbol{\beta}^\top \mathbf{A}_n \boldsymbol{\beta} + B \sum_{i=1}^n \beta_i\right),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is constant at B. θ — temperature parameter.

• In social networks, to study trends in opinions (voting choices), where G_N could be determined by "friendships" within the network.

$$\mathbb{P}(\boldsymbol{\beta}) := \frac{1}{Z_n} \exp\left(\frac{\theta}{2} \boldsymbol{\beta}^{\top} \mathbf{A}_n \boldsymbol{\beta} + B \sum_{i=1}^n \beta_i\right),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is constant at *B*. θ — temperature parameter.

- In social networks, to study trends in opinions (voting choices), where G_N could be determined by "friendships" within the network.
- (Phase transition) Modeling ferromagnetic properties (i.e., sharp change in magnetic properties of magnetic materials when heated beyond a certain (Curie) temperature coded into \mathbf{A}_n).

$$\mathbb{P}(\boldsymbol{\beta}) := \frac{1}{Z_n} \exp \left(\frac{\theta}{2} \boldsymbol{\beta}^\top \mathbf{A}_n \boldsymbol{\beta} + B \sum_{i=1}^n \beta_i \right),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is constant at B. θ — temperature parameter.

- In social networks, to study trends in opinions (voting choices), where G_N could be determined by "friendships" within the network.
- (Phase transition) Modeling ferromagnetic properties (i.e., sharp change in magnetic properties of magnetic materials when heated beyond a certain (Curie) temperature coded into \mathbf{A}_n).

Natural question

Behavior of the "sufficient statistic"

$$n^{-??} \sum_{i=1}^{n} (\beta_i - ??) \stackrel{d}{\longrightarrow} \dots$$

$$\mathbb{P}(\boldsymbol{\beta}) := \frac{1}{Z_n} \exp\left(\frac{\theta}{2} \boldsymbol{\beta}^{\top} \mathbf{A}_n \boldsymbol{\beta} + B \sum_{i=1}^n \beta_i\right),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is constant at B. θ — temperature parameter.

- In social networks, to study trends in opinions (voting choices), where G_N could be determined by "friendships" within the network.
- (Phase transition) Modeling ferromagnetic properties (i.e., sharp change in magnetic properties of magnetic materials when heated beyond a certain (Curie) temperature coded into \mathbf{A}_n).

Natural question

Behavior of the "sufficient statistic"

$$n^{-??} \sum_{i=1}^{n} (\beta_i - ??) \stackrel{d}{\longrightarrow} \dots$$

CLT for Sufficient statistic \leftrightarrow CLT for MLE of B.

$$\mathbb{P}(\boldsymbol{\beta}) := \frac{1}{Z_n} \exp\left(\frac{\theta}{2} \boldsymbol{\beta}^{\top} \mathbf{A}_n \boldsymbol{\beta} + B \sum_{i=1}^n \beta_i\right),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is constant at B. θ — temperature parameter.

- In social networks, to study trends in opinions (voting choices), where G_N could be determined by "friendships" within the network.
- (Phase transition) Modeling ferromagnetic properties (i.e., sharp change in magnetic properties of magnetic materials when heated beyond a certain (Curie) temperature coded into \mathbf{A}_n).

Natural question

Behavior of the "sufficient statistic"

$$n^{-??} \sum_{i=1}^{n} (\beta_i - ??) \stackrel{d}{\longrightarrow} \dots$$

CLT for Sufficient statistic \leftrightarrow CLT for MLE of B. What is the centering and scaling in terms of \mathbf{A}_n and B?

$$y = X\beta + \varepsilon$$
.

$$y = X\beta + \varepsilon$$
.

- Here
 - $\mathbf{y} \in \mathbb{R}^n$ is the observed data,

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}.$$

- Here
 - $\mathbf{y} \in \mathbb{R}^n$ is the observed data,
 - $\mathbf{X} \in \mathbb{R}^{n \times p}$ is the (non-random) design matrix;

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}.$$

- Here
 - $\mathbf{y} \in \mathbb{R}^n$ is the observed data,
 - $\mathbf{X} \in \mathbb{R}^{n \times p}$ is the (non-random) design matrix;
 - β is the coefficient vector;

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}.$$

- Here
 - $\mathbf{y} \in \mathbb{R}^n$ is the observed data,
 - $\mathbf{X} \in \mathbb{R}^{n \times p}$ is the (non-random) design matrix;
 - β is the coefficient vector;
 - $\varepsilon \in \mathbb{R}^n$ is a vector of i.i.d. $N(0, \sigma^2)$ random variables.

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}.$$

- Here
 - $\mathbf{y} \in \mathbb{R}^n$ is the observed data,
 - $\mathbf{X} \in \mathbb{R}^{n \times p}$ is the (non-random) design matrix;
 - β is the coefficient vector;
 - $\varepsilon \in \mathbb{R}^n$ is a vector of i.i.d. $N(0, \sigma^2)$ random variables.
- Assume that σ is known (and equals 1), but the parameter $\boldsymbol{\beta} \in [-1,1]^p$ is unknown.

• Suppose π is a probability distribution on [-1,1].

- Suppose π is a probability distribution on [-1, 1].
- Let β have a product prior, under which $\{\beta_i\}_{1 \leq i \leq p} \stackrel{i.i.d.}{\sim} \pi$.

- Suppose π is a probability distribution on [-1, 1].
- Let β have a product prior, under which $\{\beta_i\}_{1 \leq i \leq p} \stackrel{i.i.d.}{\sim} \pi$.
- Then the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ of $\boldsymbol{\beta}$ given \mathbf{y} is given by

$$\frac{d\mu}{d\pi^{\otimes p}}(\boldsymbol{\beta}) \propto \exp\left(-\frac{1}{2}\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2\right)$$

- Suppose π is a probability distribution on [-1, 1].
- Let β have a product prior, under which $\{\beta_i\}_{1 \leq i \leq p} \stackrel{i.i.d.}{\sim} \pi$.
- Then the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ of $\boldsymbol{\beta}$ given \mathbf{y} is given by

$$\frac{d\mu}{d\pi^{\otimes p}}(\boldsymbol{\beta}) \propto \exp\left(-\frac{1}{2}\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2\right)$$

• The posterior is a quadratic interaction model where interaction matrix comes from $\sigma^{-2}\mathbf{X}^{\top}\mathbf{X}$ and the field vector $\sigma^{-2}\mathbf{X}^{\top}\mathbf{y}$ is non-constant.

- Suppose π is a probability distribution on [-1, 1].
- Let β have a product prior, under which $\{\beta_i\}_{1 \leq i \leq p} \stackrel{i.i.d.}{\sim} \pi$.
- Then the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ of $\boldsymbol{\beta}$ given \mathbf{y} is given by

$$\frac{d\mu}{d\pi^{\otimes p}}(\boldsymbol{\beta}) \propto \exp\left(-\frac{1}{2}\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2\right)$$

- The posterior is a quadratic interaction model where interaction matrix comes from $\sigma^{-2}\mathbf{X}^{\top}\mathbf{X}$ and the field vector $\sigma^{-2}\mathbf{X}^{\top}\mathbf{y}$ is non-constant.
- We want to understand the behavior of this posterior distribution.

• Most of the existing literature on this problem focuses on the setting where the operator norm $\|\mathbf{X}\| \to \infty$, which we will refer to as high signal-to-noise (SNR) regime

• Most of the existing literature on this problem focuses on the setting where the operator norm $\|\mathbf{X}\| \to \infty$, which we will refer to as high signal-to-noise (SNR) regime — posterior contraction + Laplace method

- Most of the existing literature on this problem focuses on the setting where the operator norm $\|\mathbf{X}\| \to \infty$, which we will refer to as high signal-to-noise (SNR) regime posterior contraction + Laplace method
- A different regime is the case when $\|\mathbf{X}\|_2 = O(1)$, which results in the low SNR regime

- Most of the existing literature on this problem focuses on the setting where the operator norm $\|\mathbf{X}\| \to \infty$, which we will refer to as high signal-to-noise (SNR) regime posterior contraction + Laplace method
- A different regime is the case when $\|\mathbf{X}\|_2 = O(1)$, which results in the low SNR regime no posterior contraction + no Laplace method

- Most of the existing literature on this problem focuses on the setting where the operator norm $\|\mathbf{X}\| \to \infty$, which we will refer to as high signal-to-noise (SNR) regime posterior contraction + Laplace method
- A different regime is the case when $\|\mathbf{X}\|_2 = O(1)$, which results in the low SNR regime no posterior contraction + no Laplace method
- To simplify, let's focus on the very special case p = 1:

$$Y_1, \dots, Y_n \stackrel{iid}{\sim} N(\theta, 1) \text{ vs. } Y_1, \dots, Y_n \stackrel{iid}{\sim} N\left(\frac{h}{\sqrt{n}}, 1\right).$$

- Most of the existing literature on this problem focuses on the setting where the operator norm $\|\mathbf{X}\| \to \infty$, which we will refer to as high signal-to-noise (SNR) regime posterior contraction + Laplace method
- A different regime is the case when $\|\mathbf{X}\|_2 = O(1)$, which results in the low SNR regime no posterior contraction + no Laplace method
- To simplify, let's focus on the very special case p = 1:

$$Y_1, \dots, Y_n \stackrel{iid}{\sim} N(\theta, 1) \text{ vs. } Y_1, \dots, Y_n \stackrel{iid}{\sim} N\left(\frac{h}{\sqrt{n}}, 1\right).$$

• The first example has design vector $\mathbf{1}_n$, which has norm \sqrt{n} .

- Most of the existing literature on this problem focuses on the setting where the operator norm $\|\mathbf{X}\| \to \infty$, which we will refer to as high signal-to-noise (SNR) regime posterior contraction + Laplace method
- A different regime is the case when $\|\mathbf{X}\|_2 = O(1)$, which results in the low SNR regime no posterior contraction + no Laplace method
- To simplify, let's focus on the very special case p = 1:

$$Y_1, \dots, Y_n \stackrel{iid}{\sim} N(\theta, 1) \text{ vs. } Y_1, \dots, Y_n \stackrel{iid}{\sim} N\left(\frac{h}{\sqrt{n}}, 1\right).$$

- The first example has design vector $\mathbf{1}_n$, which has norm \sqrt{n} .
- Hence this is high SNR.

- Most of the existing literature on this problem focuses on the setting where the operator norm $\|\mathbf{X}\| \to \infty$, which we will refer to as high signal-to-noise (SNR) regime posterior contraction + Laplace method
- A different regime is the case when $\|\mathbf{X}\|_2 = O(1)$, which results in the low SNR regime no posterior contraction + no Laplace method
- To simplify, let's focus on the very special case p = 1:

$$Y_1, \dots, Y_n \stackrel{iid}{\sim} N(\theta, 1) \text{ vs. } Y_1, \dots, Y_n \stackrel{iid}{\sim} N\left(\frac{h}{\sqrt{n}}, 1\right).$$

- The first example has design vector $\mathbf{1}_n$, which has norm \sqrt{n} .
- Hence this is high SNR. In this case consistent estimation of θ is possible, and the posterior contracts at rate \sqrt{n} (Bernstein-von-Mises).

• The second example has design vector $n^{-1/2}\mathbf{1}_n$, which has norm 1.

- The second example has design vector $n^{-1/2}\mathbf{1}_n$, which has norm 1.
- Hence this is low SNR.

- The second example has design vector $n^{-1/2}\mathbf{1}_n$, which has norm 1.
- Hence this is low SNR. In this case consistent estimation of h is impossible, and the posterior does not contract.

- The second example has design vector $n^{-1/2}\mathbf{1}_n$, which has norm 1.
- Hence this is low SNR. In this case consistent estimation of h is impossible, and the posterior does not contract.
- In a very similar manner, it is expected that more generally when $\|\mathbf{X}\|_2 = O(1)$, consistent estimation of $\boldsymbol{\beta}$ is impossible.

- The second example has design vector $n^{-1/2}\mathbf{1}_n$, which has norm 1.
- Hence this is low SNR. In this case consistent estimation of h is impossible, and the posterior does not contract.
- In a very similar manner, it is expected that more generally when $\|\mathbf{X}\|_2 = O(1)$, consistent estimation of $\boldsymbol{\beta}$ is impossible.
- This is known in the special case when the design matrix **X** has IID Gaussian entries, given by $X_{ij} \sim N\left(0, \frac{1}{n}\right)$ (Barbier et al., IEEE-20).

- The second example has design vector $n^{-1/2}\mathbf{1}_n$, which has norm 1.
- Hence this is low SNR. In this case consistent estimation of h is impossible, and the posterior does not contract.
- In a very similar manner, it is expected that more generally when $\|\mathbf{X}\|_2 = O(1)$, consistent estimation of $\boldsymbol{\beta}$ is impossible.
- This is known in the special case when the design matrix **X** has IID Gaussian entries, given by $X_{ij} \sim N\left(0, \frac{1}{n}\right)$ (Barbier et al., IEEE-20).
- Also true in the Gaussian sequence model $Y_i = \beta_i + \epsilon_i$, where the design **X** is the identity matrix

- The second example has design vector $n^{-1/2}\mathbf{1}_n$, which has norm 1.
- Hence this is low SNR. In this case consistent estimation of *h* is impossible, and the posterior does not contract.
- In a very similar manner, it is expected that more generally when $\|\mathbf{X}\|_2 = O(1)$, consistent estimation of $\boldsymbol{\beta}$ is impossible.
- This is known in the special case when the design matrix **X** has IID Gaussian entries, given by $X_{ij} \sim N\left(0, \frac{1}{n}\right)$ (Barbier et al., IEEE-20).
- Also true in the Gaussian sequence model $Y_i = \beta_i + \epsilon_i$, where the design **X** is the identity matrix
- In this talk we will focus on this low SNR regime, which is a (non-trivial) extension of the LAN regime of classical statistics.

• As suggested in the last slide, we expect the posterior to not contract, and do not expect Bernstein von Mises type CLT to hold.

- As suggested in the last slide, we expect the posterior to not contract, and do not expect Bernstein von Mises type CLT to hold.
- At a high level, because of the scaling choice, the prior never gets washed away, but competes with the likelihood.

- As suggested in the last slide, we expect the posterior to not contract, and do not expect Bernstein von Mises type CLT to hold.
- At a high level, because of the scaling choice, the prior never gets washed away, but competes with the likelihood.
- Thus one expects the prior to show up in non-trivial ways in the asymptotics, affecting LLNs and CLTs.

- As suggested in the last slide, we expect the posterior to not contract, and do not expect Bernstein von Mises type CLT to hold.
- At a high level, because of the scaling choice, the prior never gets washed away, but competes with the likelihood.
- Thus one expects the prior to show up in non-trivial ways in the asymptotics, affecting LLNs and CLTs.

Natural question

A high-dimensional limit $p, n \to \infty$

$$\sum_{i=1}^{p} q_i(\beta_i - ??) \stackrel{d}{\longrightarrow} \dots \text{ conditioned on } \mathbf{X}, \mathbf{y}$$

where β drawn from the posterior.

- As suggested in the last slide, we expect the posterior to not contract, and do not expect Bernstein von Mises type CLT to hold.
- At a high level, because of the scaling choice, the prior never gets washed away, but competes with the likelihood.
- Thus one expects the prior to show up in non-trivial ways in the asymptotics, affecting LLNs and CLTs.

Natural question

A high-dimensional limit $p, n \to \infty$

$$\sum_{i=1}^{p} q_i(\beta_i - ??) \stackrel{d}{\longrightarrow} \dots \text{ conditioned on } \mathbf{X}, \mathbf{y}$$

where $\boldsymbol{\beta}$ drawn from the posterior. Note $\mathbf{q}^{\top}\boldsymbol{\beta}$ is the posterior prediction when the covariate is \mathbf{q} .

- As suggested in the last slide, we expect the posterior to not contract, and do not expect Bernstein von Mises type CLT to hold.
- At a high level, because of the scaling choice, the prior never gets washed away, but competes with the likelihood.
- Thus one expects the prior to show up in non-trivial ways in the asymptotics, affecting LLNs and CLTs.

Natural question

A high-dimensional limit $p,n\to\infty$

$$\sum_{i=1}^{p} q_i(\beta_i - ??) \stackrel{d}{\longrightarrow} \dots \text{ conditioned on } \mathbf{X}, \mathbf{y}$$

where $\boldsymbol{\beta}$ drawn from the posterior. Note $\mathbf{q}^{\top}\boldsymbol{\beta}$ is the posterior prediction when the covariate is \mathbf{q} .

Ideally the analysis will apply to both deterministic and random \mathbf{X} , and allows for possibly dependent entries.

• For two-spin Ising model with constant magnetization on approximately regular graphs —

- For two-spin Ising model with constant magnetization on approximately regular graphs —
 - Phase transitions in the fluctuations of $\sum_{i=1}^{n} \beta_i$ and universal limit laws (free of \mathbf{A}_n).

- For two-spin Ising model with constant magnetization on approximately regular graphs
 - Phase transitions in the fluctuations of $\sum_{i=1}^{n} \beta_i$ and universal limit laws (free of \mathbf{A}_n).
 - (Potential) non-universality and lack of phase transitions for $\sum_{i=1}^{n} q_i \beta_i$ for some other **q**.

- For two-spin Ising model with constant magnetization on approximately regular graphs —
 - Phase transitions in the fluctuations of $\sum_{i=1}^{n} \beta_i$ and universal limit laws (free of \mathbf{A}_n).
 - (Potential) non-universality and lack of phase transitions for $\sum_{i=1}^{n} q_i \beta_i$ for some other **q**.
- In Bayesian linear regression (only in high temperature)
 - Effect of prior in the centering (and limiting variance).

- For two-spin Ising model with constant magnetization on approximately regular graphs
 - Phase transitions in the fluctuations of $\sum_{i=1}^{n} \beta_i$ and universal limit laws (free of \mathbf{A}_n).
 - (Potential) non-universality and lack of phase transitions for $\sum_{i=1}^{n} q_i \beta_i$ for some other **q**.
- In Bayesian linear regression (only in high temperature)
 - Effect of prior in the centering (and limiting variance).
 - Bayesian low SNR credible sets

- For two-spin Ising model with constant magnetization on approximately regular graphs —
 - Phase transitions in the fluctuations of $\sum_{i=1}^{n} \beta_i$ and universal limit laws (free of \mathbf{A}_n).
 - (Potential) non-universality and lack of phase transitions for $\sum_{i=1}^{n} q_i \beta_i$ for some other **q**.
- In Bayesian linear regression (only in high temperature)
 - Effect of prior in the centering (and limiting variance).
 - Bayesian low SNR credible sets
- (Not in the talk) General Berry-Esseen bounds https://arxiv.org/abs/2005.00710 and https://arxiv.org/abs/2503.21152.

Outline

- Introduction
- 2 The Naive Mean-Field method
- 3 Two-spin Ising model: LLN + CLT
- 4 Bayesian Linear regression : LLN + CLT
- 6 Conclusion

• One frequently used method to "understand" the posterior μ is the naive mean field (NFM) variational approximation.

- One frequently used method to "understand" the posterior μ is the naive mean field (NFM) variational approximation.
- This is obtained by projecting μ to the space of product measures on $[-1,1]^p$, using the Kullback-Leibler divergence.

- One frequently used method to "understand" the posterior μ is the naive mean field (NFM) variational approximation.
- This is obtained by projecting μ to the space of product measures on $[-1,1]^p$, using the Kullback-Leibler divergence.

Typical result (Basak and Mukherjee, 2015)

Let P_{prod} denote the space of product measures on $[-1,1]^p$, then under the Mean-Field assumption $\|\mathbf{A}\|_F^2 = o(p)$ (Frobenius norm of interaction matrix), the following holds:

$$\inf_{\nu \in P_{\mathrm{prod}}} \mathrm{KL}(\nu | \mu_{\mathbf{y}, \mathbf{X}, \pi}) = o(p).$$

- One frequently used method to "understand" the posterior μ is the naive mean field (NFM) variational approximation.
- This is obtained by projecting μ to the space of product measures on $[-1,1]^p$, using the Kullback-Leibler divergence.

Typical result (Basak and Mukherjee, 2015)

Let P_{prod} denote the space of product measures on $[-1,1]^p$, then under the Mean-Field assumption $\|\mathbf{A}\|_F^2 = o(p)$ (Frobenius norm of interaction matrix), the following holds:

$$\inf_{\nu \in P_{\mathrm{prod}}} \mathrm{KL}(\nu | \mu_{\mathbf{y}, \mathbf{X}, \pi}) = o(p).$$

• The optimization resulting from the above projection is usually easy to compute.

- One frequently used method to "understand" the posterior μ is the naive mean field (NFM) variational approximation.
- This is obtained by projecting μ to the space of product measures on $[-1,1]^p$, using the Kullback-Leibler divergence.

Typical result (Basak and Mukherjee, 2015)

Let P_{prod} denote the space of product measures on $[-1,1]^p$, then under the Mean-Field assumption $\|\mathbf{A}\|_F^2 = o(p)$ (Frobenius norm of interaction matrix), the following holds:

$$\inf_{\nu \in P_{\text{prod}}} \text{KL}(\nu | \mu_{\mathbf{y}, \mathbf{X}, \pi}) = o(p).$$

- The optimization resulting from the above projection is usually easy to compute.
- NMF is computationally efficient (see Jain et al. 2018) as opposed to MCMC based methods, particularly in high dimensions.

• Write

$$\sigma^2 \mathbf{X}' \mathbf{X} = A + D$$
, and $\sigma^2 \mathbf{c} = \mathbf{X}' \mathbf{y}$.

• Write

$$\sigma^2 \mathbf{X}' \mathbf{X} = A + D$$
, and $\sigma^2 \mathbf{c} = \mathbf{X}' \mathbf{y}$.

• Suppose for simplicity that all diagonal entries of $\mathbf{X}'\mathbf{X}$ are the same, i.e. $D=d\mathbf{I}$.

• Write

$$\sigma^2 \mathbf{X}' \mathbf{X} = A + D$$
, and $\sigma^2 \mathbf{c} = \mathbf{X}' \mathbf{y}$.

• Suppose for simplicity that all diagonal entries of $\mathbf{X}'\mathbf{X}$ are the same, i.e. $D = d\mathbf{I}$. This assumption is not needed for most of our results, but simplifies the argument+notations.

• Write

$$\sigma^2 \mathbf{X}' \mathbf{X} = A + D$$
, and $\sigma^2 \mathbf{c} = \mathbf{X}' \mathbf{y}$.

- Suppose for simplicity that all diagonal entries of $\mathbf{X}'\mathbf{X}$ are the same, i.e. $D = d\mathbf{I}$. This assumption is not needed for most of our results, but simplifies the argument+notations.
- Define a linear+quadratic of the prior π by setting

$$\frac{d\pi_{\theta,d}}{d\pi}(w) = e^{\theta x - \frac{d}{2}w^2 - \alpha(\theta)},$$

where

$$\alpha(\theta) := \alpha_{\pi,d}(\theta) = \log \int_{[-1,1]} e^{\theta w - \frac{d}{2}w^2} d\pi(w).$$

• Write

$$\sigma^2 \mathbf{X}' \mathbf{X} = A + D$$
, and $\sigma^2 \mathbf{c} = \mathbf{X}' \mathbf{y}$.

- Suppose for simplicity that all diagonal entries of $\mathbf{X}'\mathbf{X}$ are the same, i.e. $D = d\mathbf{I}$. This assumption is not needed for most of our results, but simplifies the argument+notations.
- Define a linear+quadratic of the prior π by setting

$$\frac{d\pi_{\theta,d}}{d\pi}(w) = e^{\theta x - \frac{d}{2}w^2 - \alpha(\theta)},$$

where

$$\alpha(\theta) := \alpha_{\pi,d}(\theta) = \log \int_{[-1,1]} e^{\theta w - \frac{d}{2}w^2} d\pi(w).$$

• Then $\alpha''(\theta) = Var_{\pi_{\theta,d}}(W) > 0$, and so $\alpha'(\cdot) : \mathbb{R} \mapsto (-1,1)$ is strictly monotone.

• Write

$$\sigma^2 \mathbf{X}' \mathbf{X} = A + D$$
, and $\sigma^2 \mathbf{c} = \mathbf{X}' \mathbf{y}$.

- Suppose for simplicity that all diagonal entries of $\mathbf{X}'\mathbf{X}$ are the same, i.e. $D = d\mathbf{I}$. This assumption is not needed for most of our results, but simplifies the argument+notations.
- Define a linear+quadratic of the prior π by setting

$$\frac{d\pi_{\theta,d}}{d\pi}(w) = e^{\theta x - \frac{d}{2}w^2 - \alpha(\theta)},$$

where

$$\alpha(\theta) := \alpha_{\pi,d}(\theta) = \log \int_{[-1,1]} e^{\theta w - \frac{d}{2}w^2} d\pi(w).$$

• Then $\alpha''(\theta) = Var_{\pi_{\theta,d}}(W) > 0$, and so $\alpha'(\cdot) : \mathbb{R} \mapsto (-1,1)$ is strictly monotone. Note that $\alpha(\cdot)$ depends on the prior π .

• Write

$$\sigma^2 \mathbf{X}' \mathbf{X} = A + D$$
, and $\sigma^2 \mathbf{c} = \mathbf{X}' \mathbf{y}$.

- Suppose for simplicity that all diagonal entries of $\mathbf{X}'\mathbf{X}$ are the same, i.e. $D = d\mathbf{I}$. This assumption is not needed for most of our results, but simplifies the argument+notations.
- Define a linear+quadratic of the prior π by setting

$$\frac{d\pi_{\theta,d}}{d\pi}(w) = e^{\theta x - \frac{d}{2}w^2 - \alpha(\theta)},$$

where

$$\alpha(\theta) := \alpha_{\pi,d}(\theta) = \log \int_{[-1,1]} e^{\theta w - \frac{d}{2}w^2} d\pi(w).$$

- Then $\alpha''(\theta) = Var_{\pi_{\theta,d}}(W) > 0$, and so $\alpha'(\cdot) : \mathbb{R} \mapsto (-1,1)$ is strictly monotone. Note that $\alpha(\cdot)$ depends on the prior π .
- Let $h(\cdot): (-1,1) \mapsto \mathbb{R}$ be the inverse of the function $\alpha'(\cdot)$,

• Consider the following auxiliary optimization problem

$$\sup_{\mathbf{u}\in[-1,1]^p} \left\{ -\frac{1}{2}\mathbf{u}'A\mathbf{u} + \mathbf{c}'\mathbf{u} - \sum_{i=1}^p \mathrm{KL}(\pi_{h(u_i),d}|\pi_{0,d}) \right\}. \tag{1}$$

• Consider the following auxiliary optimization problem

$$\sup_{\mathbf{u}\in[-1,1]^p} \left\{ -\frac{1}{2}\mathbf{u}'A\mathbf{u} + \mathbf{c}'\mathbf{u} - \sum_{i=1}^p \mathrm{KL}(\pi_{h(u_i),d}|\pi_{0,d}) \right\}. \tag{1}$$

• Then any optimizer over the space of product measures is of the form

$$\prod_{i=1}^p \pi_{\theta_i^{\text{opt}},d},$$

where $\boldsymbol{\theta}^{\text{opt}}$ is chosen such that $\mathbb{E}_{\pi_{\boldsymbol{\theta}^{\text{opt}},d}}(W) = \mathbf{u}^{\text{opt}}$, and \mathbf{u}^{opt} is a global optimizer to (1). \mathbf{u}^{opt} will serve as the centering term.

• Consider the following auxiliary optimization problem

$$\sup_{\mathbf{u}\in[-1,1]^p} \left\{ -\frac{1}{2}\mathbf{u}'A\mathbf{u} + \mathbf{c}'\mathbf{u} - \sum_{i=1}^p \mathrm{KL}(\pi_{h(u_i),d}|\pi_{0,d}) \right\}. \tag{1}$$

• Then any optimizer over the space of product measures is of the form

$$\prod_{i=1}^p \pi_{\theta_i^{\mathrm{opt}},d},$$

where $\boldsymbol{\theta}^{\text{opt}}$ is chosen such that $\mathbb{E}_{\pi_{\boldsymbol{\theta}^{\text{opt}},d}}(W) = \mathbf{u}^{\text{opt}}$, and \mathbf{u}^{opt} is a global optimizer to (1). \mathbf{u}^{opt} will serve as the centering term.

ullet Thus we have reduced the optimization over the space of measures to an optimization over the space of mean vectors ${f u}$.

• Consider the following auxiliary optimization problem

$$\sup_{\mathbf{u}\in[-1,1]^p} \left\{ -\frac{1}{2}\mathbf{u}'A\mathbf{u} + \mathbf{c}'\mathbf{u} - \sum_{i=1}^p \mathrm{KL}(\pi_{h(u_i),d}|\pi_{0,d}) \right\}. \tag{1}$$

• Then any optimizer over the space of product measures is of the form

$$\prod_{i=1}^p \pi_{\theta_i^{\text{opt}}, d},$$

where $\boldsymbol{\theta}^{\text{opt}}$ is chosen such that $\mathbb{E}_{\pi_{\boldsymbol{\theta}^{\text{opt}},d}}(W) = \mathbf{u}^{\text{opt}}$, and \mathbf{u}^{opt} is a global optimizer to (1). \mathbf{u}^{opt} will serve as the centering term.

• Thus we have reduced the optimization over the space of measures to an optimization over the space of mean vectors **u**. We will refer to (1) as the mean-field prediction formula.

• Consider the following auxiliary optimization problem

$$\sup_{\mathbf{u}\in[-1,1]^p} \left\{ -\frac{1}{2}\mathbf{u}'A\mathbf{u} + \mathbf{c}'\mathbf{u} - \sum_{i=1}^p \mathrm{KL}(\pi_{h(u_i),d}|\pi_{0,d}) \right\}. \tag{1}$$

• Then any optimizer over the space of product measures is of the form

$$\prod_{i=1}^p \pi_{\theta_i^{\text{opt}}, d},$$

where $\boldsymbol{\theta}^{\text{opt}}$ is chosen such that $\mathbb{E}_{\pi_{\boldsymbol{\theta}^{\text{opt}},d}}(W) = \mathbf{u}^{\text{opt}}$, and \mathbf{u}^{opt} is a global optimizer to (1). \mathbf{u}^{opt} will serve as the centering term.

- Thus we have reduced the optimization over the space of measures to an optimization over the space of mean vectors **u**. We will refer to (1) as the mean-field prediction formula.
- Note that $\mathbf{u}^{\mathrm{opt}}$ depends on $(\mathbf{y}, \mathbf{X}, \pi)$, which we omit in the notation.

• In general the mean-field prediction formula does not have a unique solution.

- In general the mean-field prediction formula does not have a unique solution.
- In M.-Sen, JMLR-2022, Lacker-M-Yeung, IMRN-2024 we provide sufficient conditions on (\mathbf{A}, π) under which there is a unique well-separated global optimizer.

- In general the mean-field prediction formula does not have a unique solution.
- In M.-Sen, JMLR-2022, Lacker-M-Yeung, IMRN-2024 we provide sufficient conditions on (\mathbf{A}, π) under which there is a unique well-separated global optimizer.
- One sufficient condition is that $\|\mathbf{A}\|_2 \leq 1 \rho$, for some ρ (free of n, p).

- In general the mean-field prediction formula does not have a unique solution.
- In M.-Sen, JMLR-2022, Lacker-M-Yeung, IMRN-2024 we provide sufficient conditions on (\mathbf{A}, π) under which there is a unique well-separated global optimizer.
- One sufficient condition is that $\|\mathbf{A}\|_2 \leq 1 \rho$, for some ρ (free of n, p).
- This corresponds to the so called high temperature regime of statistical physics.

- In general the mean-field prediction formula does not have a unique solution.
- In M.-Sen, JMLR-2022, Lacker-M-Yeung, IMRN-2024 we provide sufficient conditions on (\mathbf{A}, π) under which there is a unique well-separated global optimizer.
- One sufficient condition is that $\|\mathbf{A}\|_2 \leq 1 \rho$, for some ρ (free of n, p).
- This corresponds to the so called high temperature regime of statistical physics.
- As a comment, typically they assume the somewhat stronger assumption $\|\mathbf{A}\|_{\infty} \leq 1 \rho$ particularly for quantitative bounds (concentration inequalities).

Outline

- Introduction
- 2 The Naive Mean-Field method
- \odot Two-spin Ising model : LLN + CLT
- 4 Bayesian Linear regression : LLN + CLT
- 6 Conclusion

Recall

$$\mathbb{P}(\boldsymbol{\beta}) := \frac{1}{Z_n} \exp\left(\frac{\theta}{2} \boldsymbol{\beta}^{\top} \mathbf{A}_n \boldsymbol{\beta} + B \sum_{i=1}^n \beta_i\right),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is **constant** at B. $\theta > 0$ — temperature parameter.

• Most of existing work analyzing $T_n = n^{-1} \sum_{i=1}^n \beta_i$ focuses exclusively on the Curie-Weiss model (see Ellis-Newman (1978), Chatterjee-Shao (2011), Shao-Zhang (2017)), where \mathbf{A}_n is the (scaled) complete graph.

Recall

$$\mathbb{P}(\boldsymbol{\beta}) := \frac{1}{Z_n} \exp\left(\frac{\theta}{2} \boldsymbol{\beta}^\top \mathbf{A}_n \boldsymbol{\beta} + B \sum_{i=1}^n \beta_i\right),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is **constant** at B. $\theta > 0$ — temperature parameter.

- Most of existing work analyzing $T_n = n^{-1} \sum_{i=1}^n \beta_i$ focuses exclusively on the Curie-Weiss model (see Ellis-Newman (1978), Chatterjee-Shao (2011), Shao-Zhang (2017)), where \mathbf{A}_n is the (scaled) complete graph.
- It is thus perhaps not surprising that (X_1, \ldots, X_n) can be expressed as an exact mixture of i.i.d. laws. This observation (implicitly) plays an important role in existing analysis

Recall

$$\mathbb{P}(\boldsymbol{\beta}) := \frac{1}{Z_n} \exp\left(\frac{\theta}{2} \boldsymbol{\beta}^\top \mathbf{A}_n \boldsymbol{\beta} + B \sum_{i=1}^n \beta_i\right),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is **constant** at B. $\theta > 0$ — temperature parameter.

- Most of existing work analyzing $T_n = n^{-1} \sum_{i=1}^n \beta_i$ focuses exclusively on the Curie-Weiss model (see Ellis-Newman (1978), Chatterjee-Shao (2011), Shao-Zhang (2017)), where \mathbf{A}_n is the (scaled) complete graph.
- It is thus perhaps not surprising that (X_1, \ldots, X_n) can be expressed as an exact mixture of i.i.d. laws. This observation (implicitly) plays an important role in existing analysis

Recall

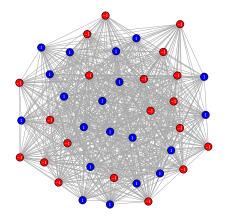
$$\mathbb{P}(\boldsymbol{\beta}) := \frac{1}{Z_n} \exp \left(\frac{\theta}{2} \boldsymbol{\beta}^\top \mathbf{A}_n \boldsymbol{\beta} + B \sum_{i=1}^n \beta_i \right),$$

 $\beta_i \in \{-1,1\}$ binary. The field vector **c** is **constant** at B. $\theta > 0$ — temperature parameter.

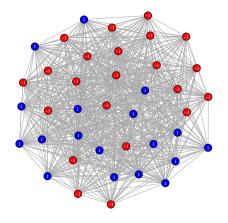
- Most of existing work analyzing $T_n = n^{-1} \sum_{i=1}^n \beta_i$ focuses exclusively on the Curie-Weiss model (see Ellis-Newman (1978), Chatterjee-Shao (2011), Shao-Zhang (2017)), where \mathbf{A}_n is the (scaled) complete graph.
- It is thus perhaps not surprising that $(X_1, ..., X_n)$ can be expressed as an exact mixture of i.i.d. laws. This observation (implicitly) plays an important role in existing analysis

We focus on other approximately regular graphs with diverging degree satisfying a Mean-Field condition.

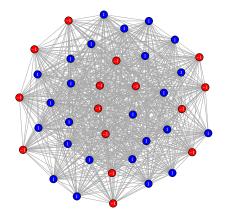
$N = 40, \ \kappa = 0.9, \ \theta = 0.5, \ B = 0$



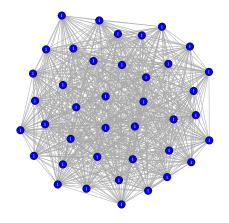
$N = 40, \ \kappa = 0.9, \ \theta = 0.5, \ B = 0$



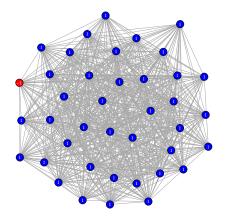
$N = 40, \ \kappa = 0.9, \ \theta = 0.5, \ B = 0$



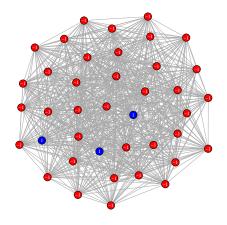
$N = 40, \ \kappa = 0.9, \ \theta = 1.5, \ B = 0$



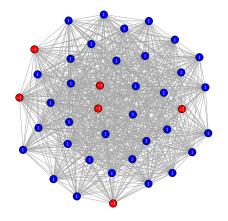
$N = 40, \ \kappa = 0.9, \ \theta = 1.5, \ B = 0$



$N = 40, \ \kappa = 0.9, \ \theta = 1.5, \ B = 0$



$N = 40, \, \kappa = 0.9, \, \theta = 1.5, \, B = 0$



$$t = \tanh(\theta t + B).$$

$$t = \tanh(\theta t + B).$$

- Properties of above equation
 - If $0 \le \theta \le 1$, B = 0, then unique solution $t_{\theta,B} = 0$.

$$t = \tanh(\theta t + B).$$

- Properties of above equation
 - If $0 \le \theta \le 1$, B = 0, then unique solution $t_{\theta,B} = 0$.
 - If $\theta > 1$, B = 0, then unique positive solution $t_{\theta,B}$.

$$t = \tanh(\theta t + B).$$

- Properties of above equation
 - If $0 \le \theta \le 1$, B = 0, then unique solution $t_{\theta,B} = 0$.
 - If $\theta > 1$, B = 0, then unique positive solution $t_{\theta,B}$.
 - If $\theta > 0$, B > 0, then unique positive solution $t_{\theta,B}$ (negative soln. for B < 0).

$$t = \tanh(\theta t + B).$$

- Properties of above equation
 - If $0 \le \theta \le 1$, B = 0, then unique solution $t_{\theta,B} = 0$.
 - If $\theta > 1$, B = 0, then unique positive solution $t_{\theta,B}$.
 - If $\theta > 0$, B > 0, then unique positive solution $t_{\theta,B}$ (negative soln. for B < 0).
- Can be extended to the case beyond two-spins as long as \mathbf{A}_n is "approximately" regular.

$$t = \tanh(\theta t + B).$$

- Properties of above equation
 - If $0 \le \theta \le 1$, B = 0, then unique solution $t_{\theta,B} = 0$.
 - If $\theta > 1$, B = 0, then unique positive solution $t_{\theta,B}$.
 - If $\theta > 0$, B > 0, then unique positive solution $t_{\theta,B}$ (negative soln. for B < 0).
- Can be extended to the case beyond two-spins as long as \mathbf{A}_n is "approximately" regular.

Does
$$n^{-??} \sum_{i=1}^{n} (\beta_i - t_{\theta,B})$$
 converge?

Basak-Mukherjee (2015), D.-Mukherjee (2023)

Basak-Mukherjee (2015), D.-Mukherjee (2023)

• If
$$B > 0$$
, then $T_n \stackrel{d}{\longrightarrow} t_{\theta,B}$

Basak-Mukherjee (2015), D.-Mukherjee (2023)

- If B > 0, then $T_n \stackrel{d}{\longrightarrow} t_{\theta,B}$

Basak-Mukherjee (2015), D.-Mukherjee (2023)

- If B > 0, then $T_n \stackrel{d}{\longrightarrow} t_{\theta,B}$
- If $\theta > 1, B = 0$ and G_n is "well-connected", then

$$T_n \xrightarrow{w} \begin{cases} t_{\theta,B} & \text{w.p. } 0.5 \\ -t_{\theta,B} & \text{w.p. } 0.5 \end{cases}$$

Basak-Mukherjee (2015), D.-Mukherjee (2023)

- If $\theta > 1, B = 0$ and G_n is "well-connected", then

$$T_n \xrightarrow{w} \begin{cases} t_{\theta,B} & \text{w.p. } 0.5 \\ -t_{\theta,B} & \text{w.p. } 0.5 \end{cases}$$

- Shades of phase transition at $\theta = 1$ when B = 0.
- More profound effect of phase transition in rates of convergence and fluctuations.

B > 0, D.-Mukherjee (2023)

If \mathbf{A}_n is a sequence of "approximately" d_n -regular graphs, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(\sqrt{n}(T_n - t_{\theta,B}) \le x) - \mathbb{P}\left(\sqrt{\frac{1 - t_{\theta,B}^2}{1 - \theta(1 - t_{\theta,B}^2)}} Z \le x\right) \right| \lesssim \frac{\sqrt{n}}{d_n}.$$

B > 0, D.-Mukherjee (2023)

If \mathbf{A}_n is a sequence of "approximately" d_n -regular graphs, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(\sqrt{n}(T_n - t_{\theta,B}) \le x) - \mathbb{P}\left(\sqrt{\frac{1 - t_{\theta,B}^2}{1 - \theta(1 - t_{\theta,B}^2)}} Z \le x\right) \right| \lesssim \frac{\sqrt{n}}{d_n}.$$

- In particular if $d_n \gg n^{1/2}$, then the distributional limit holds.
- This threshold is tight as there exists sequence of graphs with $d_n \sim n^{1/2}$ for which the limit does not hold.

$B = 0, \, \theta < 1, \, \text{D.-Mukherjee} \, (2023)$

If \mathbf{A}_n is a sequence of "approximately" d_n -regular graphs, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(\sqrt{n}T_n \le x) - \mathbb{P}\left(\sqrt{\frac{1}{1-\theta}}Z \le x\right) \right| \lesssim \frac{n^{1/3} \mathrm{poly}(\log n)}{d_n}.$$

$B = 0, \, \theta < 1, \, \text{D.-Mukherjee} \, (2023)$

If \mathbf{A}_n is a sequence of "approximately" d_n -regular graphs, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(\sqrt{n}T_n \le x) - \mathbb{P}\left(\sqrt{\frac{1}{1-\theta}}Z \le x\right) \right| \lesssim \frac{n^{1/3} \mathrm{poly}(\log n)}{d_n}.$$

- In particular if $d_n \gg n^{1/3}$, then the distributional limit holds.
- We do expect the limit to hold whenever $d_n \to \infty$.

B = 0, $\theta = 1$, D.-Mukherjee (2023)

If G_N is a sequence of "approximately" d_n -regular graphs which are well-connected, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(n^{1/4} T_n \le x) - \mathbb{P}(W \le x) \right| \lesssim \frac{\sqrt{n} \operatorname{poly}(\log n)}{d_n}$$

where W now has density proportional to $\exp(-x^4/12)$.

B = 0, $\theta = 1$, D.-Mukherjee (2023)

If G_N is a sequence of "approximately" d_n -regular graphs which are well-connected, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(n^{1/4} T_n \le x) - \mathbb{P}(W \le x) \right| \lesssim \frac{\sqrt{n} \operatorname{poly}(\log n)}{d_n}$$

where W now has density proportional to $\exp(-x^4/12)$.

- In particular if $d_n \gg n^{1/2}$, then the distributional limit holds.
- This threshold is tight as there exists sequence of graphs with $d_n \sim n^{1/2}$ for which the limit does not hold.

B = 0, $\theta = 1$, D.-Mukherjee (2023)

If G_N is a sequence of "approximately" d_n -regular graphs which are well-connected, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(n^{1/4} T_n \le x) - \mathbb{P}(W \le x) \right| \lesssim \frac{\sqrt{n} \operatorname{poly}(\log n)}{d_n}$$

where W now has density proportional to $\exp(-x^4/12)$.

- In particular if $d_n \gg n^{1/2}$, then the distributional limit holds.
- This threshold is tight as there exists sequence of graphs with $d_n \sim n^{1/2}$ for which the limit does not hold.
- The well-connectedness assumption is also tight.

$B = 0, \theta > 1$, D.-Mukherjee (2023)

If \mathbf{A}_n is a sequence of "approximately" d_n -regular graphs which are well-connected, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(\sqrt{n}(T_n - t_{\theta,B}) \le x | T_n > 0) - \mathbb{P}\left(\sqrt{\frac{1 - t_{\theta,B}^2}{1 - \theta(1 - t_{\theta,B}^2)}} Z \le x\right) \right| \lesssim \frac{\sqrt{n}}{d_n}$$

Similar result holds by conditioning on $T_n < 0$ and replacing $t_{\theta,B}$ by $-t_{\theta,B}$.

$B=0, \theta > 1, D.$ -Mukherjee (2023)

If \mathbf{A}_n is a sequence of "approximately" d_n -regular graphs which are well-connected, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(\sqrt{n}(T_n - t_{\theta,B}) \le x | T_n > 0) - \mathbb{P}\left(\sqrt{\frac{1 - t_{\theta,B}^2}{1 - \theta(1 - t_{\theta,B}^2)}} Z \le x\right) \right| \lesssim \frac{\sqrt{n}}{d_n}$$

Similar result holds by conditioning on $T_n < 0$ and replacing $t_{\theta,B}$ by $-t_{\theta,B}$.

- In particular if $d_n \gg n^{1/2}$, then the distributional limit holds.
- This threshold is tight as there exists sequence of graphs with $d_n \sim n^{1/2}$ for which the limit does not hold.

$B = 0, \theta > 1$, D.-Mukherjee (2023)

If \mathbf{A}_n is a sequence of "approximately" d_n -regular graphs which are well-connected, then:

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}(\sqrt{n}(T_n - t_{\theta,B}) \le x | T_n > 0) - \mathbb{P}\left(\sqrt{\frac{1 - t_{\theta,B}^2}{1 - \theta(1 - t_{\theta,B}^2)}} Z \le x\right) \right| \lesssim \frac{\sqrt{n}}{d_n}$$

Similar result holds by conditioning on $T_n < 0$ and replacing $t_{\theta,B}$ by $-t_{\theta,B}$.

- In particular if $d_n \gg n^{1/2}$, then the distributional limit holds.
- This threshold is tight as there exists sequence of graphs with $d_n \sim n^{1/2}$ for which the limit does not hold.
- The well-connectedness assumption is also tight.

• Lack of phase transition or universality for "approximately" d_n -regular graphs.

• Lack of phase transition or universality for "approximately" d_n -regular graphs.

Other directions in Erdős-Rényi/ random regular

Suppose \mathbf{A}_n is the adjacency matrix of an Erdős-Rényi graph with edge probability κ_n , scaled by the edge density $n\kappa_n$. Assume the following

$$\kappa_n >> n^{-1/2}, \quad \|\mathbf{q}\|_2 = 1, \quad \sum_{i=1}^n q_i = o(n^{1/2}).$$

Then

$$\sum_{i=1}^{n} q_i(\beta_i - t_{\theta,B}) | \mathbf{A}_n \xrightarrow{w} N(0, 1 - t_{\theta,B}^2).$$

• Lack of phase transition or universality for "approximately" d_n -regular graphs.

Other directions in Erdős-Rényi/ random regular

Suppose \mathbf{A}_n is the adjacency matrix of an Erdős-Rényi graph with edge probability κ_n , scaled by the edge density $n\kappa_n$. Assume the following

$$\kappa_n >> n^{-1/2}, \quad \|\mathbf{q}\|_2 = 1, \quad \sum_{i=1}^n q_i = o(n^{1/2}).$$

Then

$$\sum_{i=1}^{n} q_i(\beta_i - t_{\theta,B}) | \mathbf{A}_n \xrightarrow{w} N(0, 1 - t_{\theta,B}^2).$$

Therefore when **q** is a contrast, there is no phase transition!

• Lack of phase transition or universality for "approximately" d_n -regular graphs.

Other directions in Erdős-Rényi/ random regular

Suppose \mathbf{A}_n is the adjacency matrix of an Erdős-Rényi graph with edge probability κ_n , scaled by the edge density $n\kappa_n$. Assume the following

$$\kappa_n >> n^{-1/2}, \quad \|\mathbf{q}\|_2 = 1, \quad \sum_{i=1}^n q_i = o(n^{1/2}).$$

Then

$$\sum_{i=1}^{n} q_i(\beta_i - t_{\theta,B}) | \mathbf{A}_n \xrightarrow{w} N(0, 1 - t_{\theta,B}^2).$$

Therefore when **q** is a contrast, there is no phase transition!

• Generally it matters which eigenvalue of \mathbf{A}_n , the direction \mathbf{q} is most aligned towards. The corresponding eigenvalue shows up in the limiting variance.

What happens for $\mathbf{q}^{\top}\boldsymbol{\beta}$?

• Lack of phase transition or universality for "approximately" d_n -regular graphs.

Other directions in Erdős-Rényi/ random regular

Suppose \mathbf{A}_n is the adjacency matrix of an Erdős-Rényi graph with edge probability κ_n , scaled by the edge density $n\kappa_n$. Assume the following

$$\kappa_n >> n^{-1/2}, \quad \|\mathbf{q}\|_2 = 1, \quad \sum_{i=1}^n q_i = o(n^{1/2}).$$

Then

$$\sum_{i=1}^{n} q_i(\beta_i - t_{\theta,B}) | \mathbf{A}_n \xrightarrow{w} N(0, 1 - t_{\theta,B}^2).$$

Therefore when **q** is a contrast, there is no phase transition!

- Generally it matters which eigenvalue of \mathbf{A}_n , the direction \mathbf{q} is most aligned towards. The corresponding eigenvalue shows up in the limiting variance.
- In regular graphs, the leading eigenvector is always $n^{-1/2}\mathbf{1}$ with eigenvalue 1. Therefore choosing $\mathbf{q} = n^{-1/2}\mathbf{1}$ leads to universal behavior for "approximately" regular graphs.

• We derive Berry-Esseen bounds between $n^{-1/2} \sum_{i=1}^{n} \beta_i$ and an appropriate limit (Gaussian or otherwise) through the full Ferromagnetic parameter regime $\theta > 0$ and $B \in \mathbb{R}$.

- We derive Berry-Esseen bounds between $n^{-1/2} \sum_{i=1}^{n} \beta_i$ and an appropriate limit (Gaussian or otherwise) through the full Ferromagnetic parameter regime $\theta > 0$ and $B \in \mathbb{R}$.
- In particular, this condition holds both for deterministic and random interaction matrices.

- We derive Berry-Esseen bounds between $n^{-1/2} \sum_{i=1}^{n} \beta_i$ and an appropriate limit (Gaussian or otherwise) through the full Ferromagnetic parameter regime $\theta > 0$ and $B \in \mathbb{R}$.
- In particular, this condition holds both for deterministic and random interaction matrices.
- Also works beyond binary $\{-1,1\}$ spins. If instead the domination measure is some μ (symmetric around 0), one can still derive similar results with the phase transition point being at $(\alpha''(0))^{-1}$ where $\alpha(\cdot)$ is the cumulant generating function of μ .

- We derive Berry-Esseen bounds between $n^{-1/2} \sum_{i=1}^{n} \beta_i$ and an appropriate limit (Gaussian or otherwise) through the full Ferromagnetic parameter regime $\theta > 0$ and $B \in \mathbb{R}$.
- In particular, this condition holds both for deterministic and random interaction matrices.
- Also works beyond binary $\{-1,1\}$ spins. If instead the domination measure is some μ (symmetric around 0), one can still derive similar results with the phase transition point being at $(\alpha''(0))^{-1}$ where $\alpha(\cdot)$ is the cumulant generating function of μ .
- Non-universal behavior and lack of phase transitions for general linear combinations with q.

Outline

- Introduction
- 2 The Naive Mean-Field method
- 3 Two-spin Ising model: LLN + CLT
- 4 Bayesian Linear regression : LLN + CLT
- 6 Conclusion

Our goal is a high-dimensional limit $p,n\to\infty$

$$\sum_{i=1}^{p} q_i(\beta_i - ??) \stackrel{d}{\longrightarrow} \dots \text{ conditioned on } \mathbf{X}, \mathbf{y}$$

where $\boldsymbol{\beta}$ drawn from the posterior $\mu_{\mathbf{y}, \mathbf{X}, \pi}$.

Our goal is a high-dimensional limit $p,n\to\infty$

$$\sum_{i=1}^{p} q_i(\beta_i - ??) \stackrel{d}{\longrightarrow} \dots \text{ conditioned on } \mathbf{X}, \mathbf{y}$$

where β drawn from the posterior $\mu_{\mathbf{y},\mathbf{X},\pi}$.

• We make a high temperature assumption $\|\mathbf{A}\|_4 \leq 1 - \rho$, $\mathbf{A} = \mathrm{Off}(\mathbf{X}^{\top}\mathbf{X})$.

Our goal is a high-dimensional limit $p,n\to\infty$

$$\sum_{i=1}^{p} q_i (\beta_i - ??) \stackrel{d}{\longrightarrow} \dots \text{ conditioned on } \mathbf{X}, \mathbf{y}$$

where $\boldsymbol{\beta}$ drawn from the posterior $\mu_{\mathbf{y},\mathbf{X},\pi}$.

- We make a high temperature assumption $\|\mathbf{A}\|_4 \leq 1 \rho$, $\mathbf{A} = \mathrm{Off}(\mathbf{X}^{\top}\mathbf{X})$.
 - In particular, this assumption guarantees uniqueness of the Mean-Field optimizer $\mathbf{u}^{\mathrm{opt}}$ (which will be our centering), for any prior π .

Our goal is a high-dimensional limit $p,n\to\infty$

$$\sum_{i=1}^{p} q_i(\beta_i - ??) \stackrel{d}{\longrightarrow} \dots \text{ conditioned on } \mathbf{X}, \mathbf{y}$$

where β drawn from the posterior $\mu_{\mathbf{y},\mathbf{X},\pi}$.

- We make a high temperature assumption $\|\mathbf{A}\|_4 \leq 1 \rho$, $\mathbf{A} = \mathrm{Off}(\mathbf{X}^{\top}\mathbf{X})$.
 - In particular, this assumption guarantees uniqueness of the Mean-Field optimizer \mathbf{u}^{opt} (which will be our centering), for any prior π .
 - Also, this assumption ensures that we are in low SNR regime $\|\mathbf{A}\|_2 = O(1)$.

Our goal is a high-dimensional limit $p,n\to\infty$

$$\sum_{i=1}^{p} q_i(\beta_i - ??) \stackrel{d}{\longrightarrow} \dots \text{ conditioned on } \mathbf{X}, \mathbf{y}$$

where β drawn from the posterior $\mu_{\mathbf{y},\mathbf{X},\pi}$.

- We make a high temperature assumption $\|\mathbf{A}\|_4 \leq 1 \rho$, $\mathbf{A} = \mathrm{Off}(\mathbf{X}^{\top}\mathbf{X})$.
 - In particular, this assumption guarantees uniqueness of the Mean-Field optimizer \mathbf{u}^{opt} (which will be our centering), for any prior π .
 - Also, this assumption ensures that we are in low SNR regime $\|\mathbf{A}\|_2 = O(1)$.
- Our second assumption is the strong mean-field condition

$$\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}).$$

Our goal is a high-dimensional limit $p,n\to\infty$

$$\sum_{i=1}^{p} q_i(\beta_i - ??) \stackrel{d}{\longrightarrow} \dots \text{ conditioned on } \mathbf{X}, \mathbf{y}$$

where β drawn from the posterior $\mu_{\mathbf{y},\mathbf{X},\pi}$.

- We make a high temperature assumption $\|\mathbf{A}\|_4 \leq 1 \rho$, $\mathbf{A} = \mathrm{Off}(\mathbf{X}^{\top}\mathbf{X})$.
 - In particular, this assumption guarantees uniqueness of the Mean-Field optimizer $\mathbf{u}^{\mathrm{opt}}$ (which will be our centering), for any prior π .
 - Also, this assumption ensures that we are in low SNR regime $\|\mathbf{A}\|_2 = O(1)$.
- Our second assumption is the strong mean-field condition

$$\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}).$$

• Finally, we assume that **q** is an arbitrary vector with $\|\mathbf{q}\|_2 = 1$.

LLN

Lee-D.-Mukherjee, 25+

(i) Under the above assumptions, there is a unique well separated optimizer $\mathbf{u}^{\mathrm{opt}}$ for the mean-field prediction formula.

- (i) Under the above assumptions, there is a unique well separated optimizer $\mathbf{u}^{\mathrm{opt}}$ for the mean-field prediction formula.
- (ii) Further we have

$$\mathbb{E}_{\mu} \left[\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \right]^2 = O(1).$$

- (i) Under the above assumptions, there is a unique well separated optimizer $\mathbf{u}^{\mathrm{opt}}$ for the mean-field prediction formula.
- (ii) Further we have

$$\mathbb{E}_{\mu} \left[\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \right]^2 = O(1).$$

• Note that this bound is the best possible, as even if $\{\beta_i\}_{1 \leq i \leq p}$ were independent with mean u_i^{opt} , the second moment would be O(1).

- (i) Under the above assumptions, there is a unique well separated optimizer $\mathbf{u}^{\mathrm{opt}}$ for the mean-field prediction formula.
- (ii) Further we have

$$\mathbb{E}_{\mu} \left[\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \right]^2 = O(1).$$

- Note that this bound is the best possible, as even if $\{\beta_i\}_{1 \leq i \leq p}$ were independent with mean u_i^{opt} , the second moment would be O(1).
- Recall that Bayes optimal estimator for $\sum_{i=1}^{p} q_i \beta_i$ is $\sum_{i=1}^{p} q_i \mathbb{E}_{\mu} [\beta_i | \mathbf{y}, \mathbf{X}].$

- (i) Under the above assumptions, there is a unique well separated optimizer $\mathbf{u}^{\mathrm{opt}}$ for the mean-field prediction formula.
- (ii) Further we have

$$\mathbb{E}_{\mu} \left[\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \right]^2 = O(1).$$

- Note that this bound is the best possible, as even if $\{\beta_i\}_{1 \leq i \leq p}$ were independent with mean u_i^{opt} , the second moment would be O(1).
- Recall that Bayes optimal estimator for $\sum_{i=1}^{p} q_i \beta_i$ is $\sum_{i=1}^{p} q_i \mathbb{E}_{\mu}[\beta_i | \mathbf{y}, \mathbf{X}]$. Same computation shows that the Mean-Field estimator $\sum_{i=1}^{p} q_i u_i^{\text{opt}}$ is approximately Bayes optimal for $\sum_{i=1}^{p} q_i \beta_i$.

ullet The above theorem applies for deterministic ${f y}$ and ${f X}$.

- \bullet The above theorem applies for deterministic y and X.
- If y (and X) are random, it continues to hold, after conditioning on y (and X).

- ullet The above theorem applies for deterministic ${f y}$ and ${f X}$.
- If y (and X) are random, it continues to hold, after conditioning on y (and X).
- The crucial assumption of the theorem is the strong mean-field assumption $\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^2 = o(p^{-1/2})$.

- ullet The above theorem applies for deterministic ${f y}$ and ${f X}$.
- If y (and X) are random, it continues to hold, after conditioning on y (and X).
- The crucial assumption of the theorem is the strong mean-field assumption $\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^2 = o(p^{-1/2})$.
- There are known examples which show that the mean-field centering \mathbf{u}^{opt} is not the right one if $\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^2 = O(p^{-1/2})$.

• For CLT1 we still make the assumption that $X_{ij} = n^{-1/2}Z_{ij}$, where $(Z_{ij})_{1 \le i \le n, 1 \le j \le p}$ are IID observations from a sub-gaussian distribution.

- For CLT1 we still make the assumption that $X_{ij} = n^{-1/2}Z_{ij}$, where $(Z_{ij})_{1 \le i \le n, 1 \le j \le p}$ are IID observations from a sub-gaussian distribution.
- We also assume that the vector \mathbf{q} is delocalized, in the sense that $\|\mathbf{q}\|_{\infty} \to 0$.

- For CLT1 we still make the assumption that $X_{ij} = n^{-1/2}Z_{ij}$, where $(Z_{ij})_{1 \le i \le n, 1 \le j \le p}$ are IID observations from a sub-gaussian distribution.
- We also assume that the vector \mathbf{q} is delocalized, in the sense that $\|\mathbf{q}\|_{\infty} \to 0$.
- Note that if \mathbf{q} is not delocalized, CLT does not hold for $\mathbf{q}^{\top}\boldsymbol{\beta}$ even for IID random variables $\boldsymbol{\beta}$.

- For CLT1 we still make the assumption that $X_{ij} = n^{-1/2}Z_{ij}$, where $(Z_{ij})_{1 \le i \le n, 1 \le j \le p}$ are IID observations from a sub-gaussian distribution.
- We also assume that the vector \mathbf{q} is delocalized, in the sense that $\|\mathbf{q}\|_{\infty} \to 0$.
- Note that if \mathbf{q} is not delocalized, CLT does not hold for $\mathbf{q}^{\top}\boldsymbol{\beta}$ even for IID random variables $\boldsymbol{\beta}$.
- \bullet Finally, we also assume that the true data ${\bf y}$ is generated from a frequentist linear model of the form

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta}^{\star} + \boldsymbol{\varepsilon}$$
, where $\|\boldsymbol{\beta}^{\star}\|_{\infty} \leq 1$ and $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \mathbf{I})$.

CLT1

Lee-D.-Mukherjee, 25+

Under the above assumptions, conditioning on $\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}^{\star}$, under the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ we have

$$\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \stackrel{d}{\approx} N(0, v),$$

provided $p \ll n^{-2/3}$.

CLT1

Lee-D.-Mukherjee, 25+

Under the above assumptions, conditioning on $\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}^{\star}$, under the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ we have

$$\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \stackrel{d}{\approx} N(0, v),$$

provided $p \ll n^{-2/3}$. Here $v := \sum_{i=1}^p q_i^2 \alpha''(c_i)$. Recall $\mathbf{c} = \mathbf{X}^\top \mathbf{y}$.

Under the above assumptions, conditioning on $\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}^{\star}$, under the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ we have

$$\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \stackrel{d}{\approx} N(0, v),$$

provided $p \ll n^{-2/3}$. Here $v := \sum_{i=1}^p q_i^2 \alpha''(c_i)$. Recall $\mathbf{c} = \mathbf{X}^\top \mathbf{y}$.

• Note that all the quantities

$$\alpha(\cdot) = \alpha_{\pi,d}(\cdot), \quad \mathbf{c} = \mathbf{X}^{\mathsf{T}} \mathbf{y}, \quad \mathbf{u}^{\mathrm{opt}} = \mathbf{u}_{\mathbf{y},\mathbf{X},\pi}^{\mathrm{opt}}$$

can be computed once we know $(\mathbf{y}, \mathbf{X}, \pi)$. Here π -prior and d=Diagonal of $\mathbf{X}^{\top}\mathbf{X}$.

Under the above assumptions, conditioning on $\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}^{\star}$, under the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ we have

$$\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \stackrel{d}{\approx} N(0, v),$$

provided $p \ll n^{-2/3}$. Here $v := \sum_{i=1}^p q_i^2 \alpha''(c_i)$. Recall $\mathbf{c} = \mathbf{X}^\top \mathbf{y}$.

• Note that all the quantities

$$\alpha(\cdot) = \alpha_{\pi,d}(\cdot), \quad \mathbf{c} = \mathbf{X}^{\top} \mathbf{y}, \quad \mathbf{u}^{\text{opt}} = \mathbf{u}^{\text{opt}}_{\mathbf{y},\mathbf{X},\pi}$$

can be computed once we know $(\mathbf{y}, \mathbf{X}, \pi)$. Here π -prior and d=Diagonal of $\mathbf{X}^{\top}\mathbf{X}$.

• In particular, they don't depend on the unknown true β^* .

 Note that the previous theorem does not directly make the design assumptions

$$\|\mathbf{A}\|_{4} \le 1 - \rho \text{ and } \max_{1 \le i \le p} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}),$$

 Note that the previous theorem does not directly make the design assumptions

$$\|\mathbf{A}\|_{4} \le 1 - \rho \text{ and } \max_{1 \le i \le p} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}),$$

made for the LLN.

• We verify that both the assumptions hold if $p \ll n^{-2/3}$.

 Note that the previous theorem does not directly make the design assumptions

$$\|\mathbf{A}\|_{4} \le 1 - \rho \text{ and } \max_{1 \le i \le p} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}),$$

- We verify that both the assumptions hold if $p \ll n^{-2/3}$.
- This requires a new Chevet type inequality for bounding (4,4) operator norm of random covariance matrices. A more naive $\|\mathbf{A}\|_{\infty} \leq 1 \rho$ condition would result in a weaker threshold $p \ll n^{-1/2}$.

 Note that the previous theorem does not directly make the design assumptions

$$\|\mathbf{A}\|_{4} \le 1 - \rho \text{ and } \max_{1 \le i \le p} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}),$$

- We verify that both the assumptions hold if $p \ll n^{-2/3}$.
- This requires a new Chevet type inequality for bounding (4,4) operator norm of random covariance matrices. A more naive $\|\mathbf{A}\|_{\infty} \leq 1 \rho$ condition would result in a weaker threshold $p \ll n^{-1/2}$.
- We believe the dimension dependence is sharp, in that the mean-field centering has to be adjusted to go beyond this regime.

 Note that the previous theorem does not directly make the design assumptions

$$\|\mathbf{A}\|_{4} \le 1 - \rho \text{ and } \max_{1 \le i \le p} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}),$$

- We verify that both the assumptions hold if $p \ll n^{-2/3}$.
- This requires a new Chevet type inequality for bounding (4,4) operator norm of random covariance matrices. A more naive $\|\mathbf{A}\|_{\infty} \leq 1 \rho$ condition would result in a weaker threshold $p \ll n^{-1/2}$.
- We believe the dimension dependence is sharp, in that the mean-field centering has to be adjusted to go beyond this regime.
- As a comment, similar dimension dependence also arises for Bernstein-von-Mises type CLT approximations in the high SNR regime (see Katsevich, Arxiv-2023), where they require $p \ll n^{-1/2}$.

Assumptions2 (True Bayesian model)

 \bullet For CLT2, we work with a non-random design matrix ${\bf X}.$

Assumptions2 (True Bayesian model)

- \bullet For CLT2, we work with a non-random design matrix ${\bf X}.$
- We explicitly make the assumption $\|\mathbf{A}\|_4 \leq 1 \rho$.

- ullet For CLT2, we work with a non-random design matrix ${f X}$.
- We explicitly make the assumption $\|\mathbf{A}\|_4 \leq 1 \rho$.
- We also explicitly assume the strong mean-field condition

$$\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}).$$

- ullet For CLT2, we work with a non-random design matrix ${f X}$.
- We explicitly make the assumption $\|\mathbf{A}\|_4 \leq 1 \rho$.
- We also explicitly assume the strong mean-field condition

$$\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}).$$

• But now we assume that ${\bf q}$ is an (approximate) eigenvector of $-{\bf A}$, with eigenvalue λ .

- \bullet For CLT2, we work with a non-random design matrix ${\bf X}.$
- We explicitly make the assumption $\|\mathbf{A}\|_4 \leq 1 \rho$.
- We also explicitly assume the strong mean-field condition

$$\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}).$$

- But now we assume that \mathbf{q} is an (approximate) eigenvector of $-\mathbf{A}$, with eigenvalue λ .
- Finally, we assume that the true data **y** is generated from a linear model of the form

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta}^* + \boldsymbol{\varepsilon}$$
, where $\boldsymbol{\beta}^* \stackrel{IID}{\sim} \pi^*$, and $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \mathbf{I})$.

- \bullet For CLT2, we work with a non-random design matrix **X**.
- We explicitly make the assumption $\|\mathbf{A}\|_4 \leq 1 \rho$.
- We also explicitly assume the strong mean-field condition

$$\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}).$$

- But now we assume that \mathbf{q} is an (approximate) eigenvector of $-\mathbf{A}$, with eigenvalue λ .
- Finally, we assume that the true data y is generated from a linear model of the form

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta}^{\star} + \boldsymbol{\varepsilon}$$
, where $\boldsymbol{\beta}^{\star} \stackrel{IID}{\sim} \pi^{\star}$, and $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \mathbf{I})$.

• Here π^* is a symmetric distribution with finite fourth moment.

- \bullet For CLT2, we work with a non-random design matrix **X**.
- We explicitly make the assumption $\|\mathbf{A}\|_4 \leq 1 \rho$.
- We also explicitly assume the strong mean-field condition

$$\max_{i \in [p]} \sum_{j=1}^{p} A_{ij}^{2} = o(p^{-1/2}).$$

- But now we assume that \mathbf{q} is an (approximate) eigenvector of $-\mathbf{A}$, with eigenvalue λ .
- Finally, we assume that the true data **y** is generated from a linear model of the form

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta}^{\star} + \boldsymbol{\varepsilon}$$
, where $\boldsymbol{\beta}^{\star} \stackrel{IID}{\sim} \pi^{\star}$, and $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \mathbf{I})$.

• Here π^* is a symmetric distribution with finite fourth moment. Thus we allow the prior to be misspecified as π , where the true data generating prior is π^* .

CLT2

Lee-D.-Mukherjee, 25+

Under the above assumptions, conditioning on $\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}^{\star}$, under the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ we have

$$\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \stackrel{d}{\approx} N \left(0, \frac{\upsilon}{1 - \lambda \upsilon} \right)$$

CLT2

Lee-D.-Mukherjee, 25+

Under the above assumptions, conditioning on $\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}^*$, under the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ we have

$$\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \stackrel{d}{\approx} N \left(0, \frac{\upsilon}{1 - \lambda \upsilon} \right)$$

Here $v = \sum_{i=1}^{p} q_i^2 \alpha''(c_i)$ as before. Recall $\mathbf{c} = \mathbf{X}^{\top} \mathbf{y}$.

Lee-D.-Mukherjee, 25+

Under the above assumptions, conditioning on $\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}^{\star}$, under the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ we have

$$\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \stackrel{d}{\approx} N \left(0, \frac{\upsilon}{1 - \lambda \upsilon} \right)$$

Here $v = \sum_{i=1}^{p} q_i^2 \alpha''(c_i)$ as before. Recall $\mathbf{c} = \mathbf{X}^{\top} \mathbf{y}$.

• Note that the centering or scaling quantities \mathbf{u}^{opt} and v do not depend on the (possibly unknown) true prior π^* .

Lee-D.-Mukherjee, 25+

Under the above assumptions, conditioning on $\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}^{\star}$, under the posterior $\mu = \mu_{\mathbf{y}, \mathbf{X}, \pi}$ we have

$$\sum_{i=1}^{p} q_i (\beta_i - u_i^{\text{opt}}) \stackrel{d}{\approx} N \left(0, \frac{\upsilon}{1 - \lambda \upsilon} \right)$$

Here $v = \sum_{i=1}^{p} q_i^2 \alpha''(c_i)$ as before. Recall $\mathbf{c} = \mathbf{X}^{\top} \mathbf{y}$.

- Note that the centering or scaling quantities \mathbf{u}^{opt} and v do not depend on the (possibly unknown) true prior π^* .
- For both CLTs we provide explicit convergence rates for the above theorem in Kolmogorov-Smirnov distance.

• Based on CLT2, the set

$$\mathcal{I} := \mathcal{I}(\mathbf{y}, \mathbf{X}, \pi) := \left[\sum_{i=1}^{p} q_i u_i^{\text{opt}} \pm z_{\alpha/2} \sqrt{\frac{v}{1 - \lambda v}} \right]$$

is asymptotically a $1-\alpha$ credible interval under the posterior $\mu_{\mathbf{y},\mathbf{X},\pi}.$

• Based on CLT2, the set

$$\mathcal{I} := \mathcal{I}(\mathbf{y}, \mathbf{X}, \pi) := \left[\sum_{i=1}^{p} q_i u_i^{\text{opt}} \pm z_{\alpha/2} \sqrt{\frac{v}{1 - \lambda v}} \right]$$

is asymptotically a $1-\alpha$ credible interval under the posterior $\mu_{\mathbf{y},\mathbf{x},\pi}.$

Lee-D.-Mukherjee, 25+

• Suppose we are in the set up of CLT2.

• Based on CLT2, the set

$$\mathcal{I} := \mathcal{I}(\mathbf{y}, \mathbf{X}, \pi) := \left[\sum_{i=1}^{p} q_i u_i^{\text{opt}} \pm z_{\alpha/2} \sqrt{\frac{v}{1 - \lambda v}} \right]$$

is asymptotically a $1-\alpha$ credible interval under the posterior $\mu_{\mathbf{y},\mathbf{x},\pi}$.

Lee-D.-Mukherjee, 25+

- Suppose we are in the set up of CLT2.
- Then we have

$$\mathbb{P}_{\pi^*} \Big(\sum_{i=1}^p q_i \beta_i^* \in \mathcal{I}(\mathbf{y}) \Big| \mathbf{X} \Big) \stackrel{P}{\to} F(d, \alpha, \pi, \pi^*).$$

• Based on CLT2, the set

$$\mathcal{I} := \mathcal{I}(\mathbf{y}, \mathbf{X}, \pi) := \left[\sum_{i=1}^{p} q_i u_i^{\text{opt}} \pm z_{\alpha/2} \sqrt{\frac{\upsilon}{1 - \lambda \upsilon}} \right]$$

is asymptotically a $1-\alpha$ credible interval under the posterior $\mu_{\mathbf{y},\mathbf{X},\pi}$.

Lee-D.-Mukherjee, 25+

- Suppose we are in the set up of CLT2.
- Then we have

$$\mathbb{P}_{\pi^*} \Big(\sum_{i=1}^p q_i \beta_i^* \in \mathcal{I}(\mathbf{y}) \Big| \mathbf{X} \Big) \stackrel{P}{\to} F(d, \alpha, \pi, \pi^*).$$

• In particular $F(d, \alpha, \pi^*, \pi^*) = 1 - \alpha$, so if we use the correct prior, we have asymptotically valid credible intervals.

• Based on CLT2, the set

$$\mathcal{I} := \mathcal{I}(\mathbf{y}, \mathbf{X}, \pi) := \left[\sum_{i=1}^{p} q_i u_i^{\text{opt}} \pm z_{\alpha/2} \sqrt{\frac{v}{1 - \lambda v}} \right]$$

is asymptotically a $1-\alpha$ credible interval under the posterior $\mu_{\mathbf{y},\mathbf{x},\pi}$.

Lee-D.-Mukherjee, 25+

- Suppose we are in the set up of CLT2.
- Then we have

$$\mathbb{P}_{\pi^*} \left(\sum_{i=1}^p q_i \beta_i^* \in \mathcal{I}(\mathbf{y}) \middle| \mathbf{X} \right) \stackrel{P}{\to} F(d, \alpha, \pi, \pi^*).$$

- In particular $F(d, \alpha, \pi^*, \pi^*) = 1 \alpha$, so if we use the correct prior, we have asymptotically valid credible intervals.
- Also works if you sample split to estimate π^* by some $\hat{\pi}_n$ and use it as a plug-in prior.

Outline

- Introduction
- 2 The Naive Mean-Field method
- 3 Two-spin Ising model: LLN + CLT
- 4 Bayesian Linear regression : LLN + CLT
- 6 Conclusion

ullet We give general conditions on the design matrix ${f X}$ for LLN and CLT to hold, for Bayesian Linear Regression with a product prior.

- We give general conditions on the design matrix X for LLN and CLT to hold, for Bayesian Linear Regression with a product prior.
- In particular, this condition holds both for deterministic and random matrices, and allows for dependence of entries.

- We give general conditions on the design matrix X for LLN and CLT to hold, for Bayesian Linear Regression with a product prior.
- In particular, this condition holds both for deterministic and random matrices, and allows for dependence of entries.
- We give explicit error rates in terms of the Kolmogorov-Smirnov distance.

- We give general conditions on the design matrix X for LLN and CLT to hold, for Bayesian Linear Regression with a product prior.
- In particular, this condition holds both for deterministic and random matrices, and allows for dependence of entries.
- We give explicit error rates in terms of the Kolmogorov-Smirnov distance.
- We apply our results to construct credible intervals, and compute their asymptotic coverage under possible prior misspecification.

• A natural question is whether we can study the low temperature regime, where $\|\mathbf{A}\|_4 > 1$.

• A natural question is whether we can study the low temperature regime, where $\|\mathbf{A}\|_4 > 1$. In this case one can have more than one mean-field optimizer, and the right approximation should be a mixture of product measures.

- A natural question is whether we can study the low temperature regime, where $\|\mathbf{A}\|_4 > 1$. In this case one can have more than one mean-field optimizer, and the right approximation should be a mixture of product measures.
- Another possible direction is to remove the compactness assumption, and allow for a broader class of priors.

- A natural question is whether we can study the low temperature regime, where $\|\mathbf{A}\|_4 > 1$. In this case one can have more than one mean-field optimizer, and the right approximation should be a mixture of product measures.
- Another possible direction is to remove the compactness assumption, and allow for a broader class of priors. One can assume log-concavity of the prior, which also guarantees uniqueness of optimizer.

- A natural question is whether we can study the low temperature regime, where $\|\mathbf{A}\|_4 > 1$. In this case one can have more than one mean-field optimizer, and the right approximation should be a mixture of product measures.
- Another possible direction is to remove the compactness assumption, and allow for a broader class of priors. One can assume log-concavity of the prior, which also guarantees uniqueness of optimizer.
- Finally, it remains to study similar questions for non-quadratic posteriors.

- A natural question is whether we can study the low temperature regime, where $\|\mathbf{A}\|_4 > 1$. In this case one can have more than one mean-field optimizer, and the right approximation should be a mixture of product measures.
- Another possible direction is to remove the compactness assumption, and allow for a broader class of priors. One can assume log-concavity of the prior, which also guarantees uniqueness of optimizer.
- Finally, it remains to study similar questions for non-quadratic posteriors. People apply NMF to a host of problems (GLMs, Topic Modeling).

Thank you. Questions?